

# Lecture 5 Data processing inequality and Fano inequality

September 13th, 2022

# Outline

- 1 Exercises Review
- 2 Data processing inequality
- 3 Fano's inequality
- 4 Another inequality relating probability of error

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## Example

*A  $(7, 4)$  Hamming code can correct any one error; might there be a  $(14, 8)$  code that can correct any two errors?*

## Proof.

When the decoder receives  $\mathbf{r} = \mathbf{t} + \mathbf{n}$ , his aim is to deduce both  $\mathbf{t}$  and  $\mathbf{n}$  from  $\mathbf{r}$ . If it is the case that the sender can select any transmission  $\mathbf{t}$  from a code of size  $S_{\mathbf{t}}$ , and the channel can select any noise vector from a set of size  $S_{\mathbf{n}}$ , and those two selections can be recovered from the received bit string  $\mathbf{r}$ , which is one of at most  $2^N$  possible strings, then it must be the case that

$$S_{\mathbf{t}}S_{\mathbf{n}} \leq 2^N.$$

So, for a  $(N, K)$  two-error-correcting code,

$$2^K \left[ \binom{N}{2} + \binom{N}{1} + \binom{N}{0} \right] \leq 2^N.$$

however the inequality does not hold for  $K = 8$  and  $N = 14$ , which rules out the possibility that there is a  $(14, 8)$  code that is 2-error correcting.  $\square$

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# Data processing inequality

## Lemma

If  $X \rightarrow Y \rightarrow Z$ , then

$$I(X; Y) \geq I(X; Z).$$

## Proof.

By the chain rule, we can expand mutual information in two different ways:

$$\begin{aligned} I(X; Y, Z) &= I(X; Z) + I(X; Y|Z) \\ &= I(X; Y) + I(X; Z|Y). \end{aligned}$$

Since  $X$  and  $Z$  are conditionally independent given  $Y$ , we have  $I(X; Z|Y) = 0$ . Since  $I(X; Y|Z) \geq 0$ , we have

$$I(X; Y) \geq I(X; Z).$$

The equality holds if and only if  $I(X; Y|Z) = 0$  (i.e.,  $X \rightarrow Z \rightarrow Y$  forms a Markov chain). Similarly, one can prove that  $I(Y; Z) \geq I(X; Z)$ . □



## Corollary

*If  $Z = g(Y)$ , then  $I(X; Y) \geq I(X; g(Y))$ .*

## Corollary

*If  $X \rightarrow Y \rightarrow Z$ , then  $I(X; Y|Z) \leq I(X; Y)$ .*

Note that it is also possible that  $I(X; Y|Z) > I(X; Y)$  when  $X$ ,  $Y$  and  $Z$  do not form a Markov chain.

For example, let  $X$  and  $Y$  be independent fair binary random variables, and let  $Z = X + Y$ . Then  $I(X; Y) = 0$ , but  $I(X; Y|Z) = H(X|Z) - H(X|Y, Z) = H(X|Z) = P(Z = 1)P(X|Z = 1) = \frac{1}{2}$  bit.

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# Fano's inequality

## Theorem

Let  $X$  and  $Y$  be two random variables, correlated in general. with alphabet  $\mathcal{X}$  and  $\mathcal{Y}$ , respectively, where  $\mathcal{X}$  is finite but  $\mathcal{Y}$  can be countably infinite. Let  $\hat{X} := g(Y)$  be an estimate of  $X$  from observing  $Y$ , where  $g : \mathcal{Y} \rightarrow \mathcal{X}$  is a given estimation function. Define the probability of error as

$$P_e := \Pr[\hat{X} \neq X].$$

Then the following inequality holds

$$H(X|Y) \leq h_b(P_e) + P_e \cdot \log_2(|\mathcal{X}| - 1),$$

where  $h_b(x) := -x \log_2 x - (1 - x) \log_2(1 - x)$  for  $0 \leq x \leq 1$  is the binary entropy function.

## Proof.

Define a new random variable,

$$E := \begin{cases} 1, & \text{if } g(Y) \neq X \\ 0, & \text{if } g(Y) = X. \end{cases}$$

Then using the chain rule for conditional entropy, we obtain

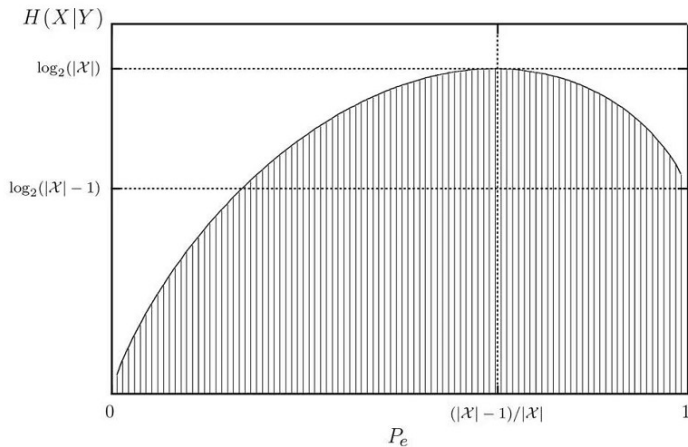
$$\begin{aligned} H(E, X|Y) &= H(X|Y) + H(E|X, Y) \\ &= H(E|Y) + H(X|E, Y). \end{aligned}$$

Observe that  $E$  is a function of  $X$  and  $Y$ ; hence,  $H(E|X, Y) = 0$ . Since conditioning never increases entropy,  $H(E|Y) \leq H(E) = h_b(P_e)$ . The remaining term,  $H(X|E, Y)$ , can be bounded as follows:

$$\begin{aligned} H(E, X|Y) &= Pr[E = 0]H(X|Y, E = 0) + Pr[E = 1]H(X|Y, E = 1) \\ &\leq (1 - P_e) \cdot 0 + P_e \cdot \log_2(|\mathcal{X}| - 1), \end{aligned}$$

since  $X = g(Y)$  for  $E = 0$ , and given  $E = 1$ , we can upper bound the conditional entropy by the logarithm of the number of remaining outcomes, i.e.,  $(|\mathcal{X}| - 1)$ . Combining these results completes the proof. □

# Permissible $(P_e, H(X|Y))$ region due to Fano's inequality



Fano's inequality yields upper and lower bounds on  $P_e$  in terms of  $H(X|Y)$ . This is illustrated in last page, where we plot the region for the pairs  $(P_e, H(X|Y))$  that are permissible under Fano's inequality.

In the figure, the boundary of the permissible (dashed) region is given by the function

$$f(P_e) := h_b(P_e) + P_e \cdot \log_2(|\mathcal{X}| - 1).$$

We obtain that when

$$\log_2(|\mathcal{X}| - 1) \leq H(X|Y) \leq \log_2(|\mathcal{X}|),$$

$P_e$  can be upper and lower bounded as follows:

$$0 < \inf\{a : f(a) \geq H(X|Y)\} \leq P_e \leq \sup\{a : f(a) \geq H(X|Y)\} < 1.$$

Furthermore, when

$$0 < H(X|Y) \leq \log_2(|\mathcal{X}| - 1),$$

only the lower bound holds:

$$P_e \geq \inf\{a : f(a) \geq H(X|Y)\} > 0.$$

Thus for all nonzero values of  $H(X|Y)$ , we obtain a lower bound (of the same form above) on  $P_e$ ; the bound implies that if  $H(X|Y)$  is bounded away from zero,  $P_e$  is also bounded away from zero.



Fano's inequality cannot be improved in the sense that the lower bound,  $H(X|Y)$ , can be achieved for some specific cases. Any bound that can be achieved in some cases is often referred to as sharp.

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From the proof of the above lemma, we can observe that equality holds in Fano's inequality, if  $H(E|Y) = H(E)$  and  $H(X|Y, E = 1) = \log_2(|\mathcal{X} - 1)$ . The former is equivalent to  $E$  being independent of  $Y$ , and the latter holds iff  $P_{X|Y}(\cdot|y)$  is uniformly distributed over the set  $\mathcal{X} \setminus \{g(y)\}$ . We can therefore create an example in which equality holds in Fano's inequality.

## Example

Suppose that  $X$  and  $Y$  are two independent random variables which are both uniformly distributed on the alphabet  $\{0, 1, 2\}$ . Let the estimating function be given by  $g(y) = y$ . Then

$$P_e = \Pr[g(Y) \neq X] = \Pr[Y \neq X] = 1 - \sum_{x=0}^2 P_X(x)P_Y(x) = \frac{2}{3}.$$

In this case, equality is achieved in Fano's inequality, i.e.,

$$h_b\left(\frac{2}{3}\right) + \frac{2}{3} \cdot \log_2(3-1) = H(X|Y) = H(X) = \log_2 3.$$

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Let  $X$  and  $X'$  be two independent identically distributed random variables with entropy  $H(X)$ . The probability at  $X = X'$  is given by

$$Pr(X = X') = \sum_x p^2(x).$$

## Lemma

*If  $X$  and  $X'$  are i.i.d. with entropy  $H(X)$ .*

$$Pr(X = X') \geq 2^{-H(X)},$$

*with equality if and only if  $X$  has a uniform distribution.*

Proof.

Suppose that  $X \sim p(x)$ . By Jensen's inequality, we have

$$2^{E \log p(x)} \leq E 2^{\log p(x)},$$

which implies that

$$2^{-H(X)} = 2^{\sum p(x) \log p(x)} \leq \sum p(x) 2^{\log p(x)} = \sum p^2(x).$$



## Corollary

Let  $X, X'$  be independent with  $X \sim p(x)$ ,  $X' \sim r(x)$ ,  $x, x' \in \mathcal{X}$ .  
Then

$$\begin{aligned} P(X = X') &\geq 2^{-H(p) - D(p||r)}, \\ P(X = X') &\geq 2^{-H(r) - D(r||p)}. \end{aligned}$$



## Proof.

We have

$$\begin{aligned} 2^{-H(p)-D(p||r)} &= 2^{\sum p(x) \log p(x) + \sum p(x) \log \frac{r(x)}{p(x)}} \\ &= 2^{\sum p(x) \log r(x)} \\ &\leq \sum p(x) 2^{\log r(x)} \\ &= \sum p(x) r(x) \\ &= P(X = X'), \end{aligned}$$

where the inequality follows from Jensen's inequality and the convexity of the function  $f(y) = 2^y$ . □